

# 11.6.2 Change of Measure for a Compound Poisson Process

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Let  $N(t)$  be a Poisson process with intensity  $\lambda$ , and let  $Y_1, Y_2, \dots$  be a sequence of identically distributed random variables defined on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . We assume the random variables  $Y_1, Y_2, \dots$  are independent of one another and also independent of the Poisson process  $N(t)$ . We define the *compound Poisson process*

$$\text{總共移動幅度 } \underline{Q(t)} = \sum_{i=1}^{N(t)} Y_i. \quad (11.6.9)$$

Note for future reference that if  $N$  jumps at time  $t$ , then  $Q$  jumps at time  $t$  and

$$Q(t) - Q(t-) = \Delta Q(t) = Y_{N(t)}. \quad (11.6.10)$$

Jumps in  $Q(t)$  **are random sizes**

First jump size:  $Y_1$

Second jump size:  $Y_2$

$Q(t)$ : value of  $Q$  immediately **after** the jump

$Q(t-)$ : value of  $Q$  immediately **before** the jump

For a Poisson process, the change of measure affects the intensity.  
For a compound Poisson process, the change of measure can affect **both** the intensity and the distribution of jump sizes.

Our goal is to change the measure so that the intensity of  $N(t)$  and the distribution of the jump sizes  $Y_1, Y_2, \dots$  both change. We first consider the case when the jump-size random variables have a discrete distribution (i.e., each  $Y_i$  takes one of finitely many possible nonzero values  $y_1, y_2, \dots, y_M$ ). Let  $p(y_m)$  denote the probability that a jump is of size  $y_m$ :

定義M種jump size的discrete distribution

$$p(y_m) = \mathbb{P}\{Y_i = y_m\}, \quad m = 1, \dots, M.$$

This does not depend on  $i$  since  $Y_1, Y_2, \dots$  are identically distributed. We assume that  $p(y_m) > 0$  for every  $m$  and, of course, that  $\sum_{m=1}^M p(y_m) = 1$ .

Let  $N_m(t)$  denote the number of jumps in  $Q(t)$  of size  $y_m$  up to and including time  $t$ , so that

$$N(t) = \sum_{m=1}^M N_m(t) \text{ and } Q(t) = \sum_{m=1}^M y_m N_m(t).$$

$N(t)$ : 跳的總次數  
第  $m$  種 jump size 跳了  $N_m(t)$  次

According to Corollary 11.3.4,  $N_1, \dots, N_M$  are independent Poisson processes and each  $N_m$  has intensity  $\lambda_m = \lambda p(y_m)$ .

Let  $\tilde{\lambda}_1, \dots, \tilde{\lambda}_M$  be given positive numbers, and set

$$Z_m(t) = e^{(\lambda_m - \tilde{\lambda}_m)t} \left( \frac{\tilde{\lambda}_m}{\lambda_m} \right)^{N_m(t)} \text{ and } Z(t) = \prod_{m=1}^M Z_m(t). \quad (11.6.11)$$

進行測度轉換要用到的  $Z$

$$\tilde{\mathbb{P}}(A) = \int_A Z(T) dP \text{ for all } Z \in \mathcal{F}.$$

$$Z_m(t) = e^{(\lambda_m - \tilde{\lambda}_m)t} \left( \frac{\tilde{\lambda}_m}{\lambda_m} \right)^{N_m(t)} \quad \text{and} \quad Z(t) = \prod_{m=1}^M Z_m(t). \quad (11.6.11)$$

**Lemma 11.6.4.** *The process  $Z(t)$  of (11.6.11) is a martingale. In particular,  $\mathbb{E}Z(t) = 1$  for all  $t$ .*

PROOF: From Lemma 11.6.1, we have

$$dZ_m(t) = \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} Z_m(t-) dM_m(t), \quad (11.6.12)$$

where

$$M_m(t) = N_m(t) - \lambda_m dt.$$

$$Z_m(t) = e^{(\lambda_m - \tilde{\lambda}_m)t} \left(\frac{\tilde{\lambda}_m}{\lambda_m}\right)^{N_m(t)} \quad \text{and} \quad Z(t) = \prod_{m=1}^M Z_m(t). \quad (11.6.11)$$

$$f(X(t)) = f(X(0)) + \int_0^t f'(X(s)) dX^c(s) + \frac{1}{2} \int_0^t f''(X(s)) dX^c(s) dX^c(s) + \sum_{0 < s \leq t} [f(X(s)) - f(X(s-))]. \quad (11.5.4)$$

$$f(x) = e^x, \quad Z_m(t) = f(X(t)), \quad X(t) = \underbrace{(\lambda_m - \tilde{\lambda}_m)t}_{X^c(t)} + \underbrace{N_m(t)}_{J(t)} \log\left(\frac{\tilde{\lambda}_m}{\lambda_m}\right)$$

$$f'(X(t)) = f(X(t)) = Z_m(t)$$

$$dX^c(t) = (\lambda_m - \tilde{\lambda}_m) dt$$

$$\text{By 11.5.4, } Z_m(t) = f(X(t)) = f(X(0)) + (\lambda_m - \tilde{\lambda}_m) \int_0^t f'(X(u)) du + \sum_{0 < u \leq t} [f(X(u)) - f(X(u-))]$$

$$= Z_m(0) + (\lambda_m - \tilde{\lambda}_m) \int_0^t Z_m(u) du + \sum_{0 < u \leq t} [Z_m(u) - Z_m(u-)]$$

$$\textcircled{1} \text{ jump at time } u = Z_m(u) = \left(\frac{\tilde{\lambda}_m}{\lambda_m}\right) Z_m(u-) \Rightarrow Z_m(u) - Z_m(u-) = \left(\frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m}\right) Z_m(u-)$$

$$\textcircled{2} \text{ no jump at time } u = Z_m(u) - Z_m(u-) = 0 \Rightarrow Z_m(u) - Z_m(u-) = \left(\frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m}\right) Z_m(u-) \underbrace{N_m(u)}_{=0 \text{ if no jump}}$$

$$\therefore \sum_{0 < u \leq t} [Z_m(u) - Z_m(u-)] = \sum_{0 < u \leq t} \left(\frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m}\right) Z_m(u-) \Delta N_m(u) = \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} \int_0^t Z_m(u-) dN_m(u)$$

$$\therefore Z_m(t) = Z_m(0) + (\lambda_m - \tilde{\lambda}_m) \int_0^t Z_m(u-) du + \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} \int_0^t Z_m(u-) dN_m(u)$$

$$= Z_m(0) - \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} \int_0^t Z_m(u-) \lambda_m du + \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} \int_0^t Z_m(u-) dN_m(u)$$

$$= Z_m(0) + \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} \int_0^t Z_m(u-) dM_m(u) \quad M_m(t) = N_m(t) - \lambda_m t \Rightarrow dN_m(t) = dM_m(t) + \lambda_m dt$$

$$\Rightarrow dZ_m(t) = \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} Z_m(t-) dM_m(t)$$

$$Z_m(t) = Z_m(0) + \frac{\tilde{\lambda}_m - \lambda_m}{\lambda_m} \int_0^t Z_m(u-) dM_m(t)$$

Because the integrand in (11.6.12) is left-continuous and the compensated Poisson process is a martingale, the process  $Z_m$  is a martingale (Theorem 11.4.5).

**Theorem 11.4.5.** Assume that the jump process  $X(s)$  of (11.4.1)–(11.4.3) is a martingale, the integrand  $\Phi(s)$  is left-continuous and adapted, and

$$\mathbb{E} \int_0^t \Gamma^2(s) \Phi^2(s) ds < \infty \text{ for all } t \geq 0.$$

Then the stochastic integral  $\int_0^t \Phi(s) dX(s)$  is also a martingale.

$\left\{ \begin{array}{l} Z_m(u-) \text{ is left continuous} \\ M_m(t) \text{ is martingale} \end{array} \right. \rightarrow Z_m \text{ is martingale}$

$\therefore Z_m(t)$  is martingale,  $Z_m(0) = 1$

$\therefore EZ_m(t) = 1$

**Corollary 11.5.5 (Itô's product rule for jump processes).** Let  $X_1(t)$  and  $X_2(t)$  be jump processes. Then

$$\begin{aligned} X_1(t)X_2(t) &= X_1(0)X_2(0) + \int_0^t X_2(s) dX_1^c(s) + \int_0^t X_1(s) dX_2^c(s) \\ &\quad + [X_1^c, X_2^c](t) + \sum_{0 < s \leq t} [X_1(s)X_2(s) - X_1(s-)X_2(s-)] \\ &= X_1(0)X_2(0) + \int_0^t X_2(s-) dX_1(s) + \int_0^t X_1(s-) dX_2(s) \\ &\quad + [X_1, X_2](t). \end{aligned} \tag{11.5.11}$$

For  $m \neq n$ , the Poisson processes  $N_m$  and  $N_n$  have no simultaneous jumps, and hence  $[Z_m, Z_n] = 0$ . Itô's product rule (Corollary 11.5.5) implies that

$$d(Z_1(t)Z_2(t)) = Z_2(t-) dZ_1(t) + Z_1(t-) dZ_2(t). \tag{11.6.13}$$

Because both  $Z_1$  and  $Z_2$  are martingales and the integrands in (11.6.13) are left-continuous, the process  $Z_1 Z_2$  is a martingale. Because  $Z_1 Z_2$  has no jumps simultaneous with the jumps of  $Z_3$ , Itô's product rule further implies

$$d(Z_1(t)Z_2(t)Z_3(t)) = Z_3(t-) d(Z_1(t)Z_2(t)) + (Z_1(t-)Z_2(t-)) dZ_3(t).$$

martingale (By Thm 11.4.5)

$$\underline{Z_1(t)Z_2(t)} = Z_1(0)Z_2(0) + \int_0^t \underline{Z_2(u-)} \underline{dZ_1(u)} + \int_0^t \underline{Z_1(u-)} \underline{dZ_2(u)} + \underline{[Z_m, Z_n](t)} \stackrel{=0}{}$$

$$\Rightarrow dZ_1(t)Z_2(t) = Z_2(t-)dZ_1(t) + Z_1(t-)dZ_2(t)$$

Once again, the integrators are martingales and the integrands are left-continuous. Therefore,  $Z_1 Z_2 Z_3$  is a martingale. Continuing this process, we eventually conclude that  $Z(t) = Z_1(t) Z_2(t) \cdots Z_m(t)$  is a martingale.  $\square$

Fix  $T > 0$ . Because  $Z(T) > 0$  almost surely and  $\mathbb{E}Z(T) = 1$ , we can use  $Z(T)$  to change the measure, defining

$$\tilde{\mathbb{P}}(A) = \int_A Z(T) dP \text{ for all } Z \in \mathcal{F}.$$

**Theorem 11.6.5 (Change of compound Poisson intensity and jump distribution for finitely many jump sizes).** Under  $\tilde{\mathbb{P}}$ ,  $Q(t)$  is a compound Poisson process with intensity  $\tilde{\lambda} = \sum_{m=1}^M \tilde{\lambda}_m$ , and  $Y_1, Y_2, \dots$  are independent, identically distributed random variables with

$$\tilde{\mathbb{P}}\{Y_i = y_m\} = \tilde{p}(y_m) = \frac{\tilde{\lambda}_m}{\tilde{\lambda}}. \quad (11.6.14)$$

**KEY STEP IN PROOF:** We use the independence of  $N_1, \dots, N_M$  under  $\mathbb{P}$  to compute the moment-generating function of  $Q(t)$  under  $\tilde{\mathbb{P}}$ . For  $0 \leq t \leq T$ , Lemma 5.2.1 and the moment-generating function formula (11.3.4) imply

$$\begin{aligned}
 \tilde{\mathbb{E}} \left[ e^{uQ(t)} \right] &\stackrel{\textcircled{1}}{=} \mathbb{E} \left[ e^{uQ(t)} Z(t) \right] \\
 &\stackrel{\textcircled{2}}{=} \mathbb{E} \left[ \exp \left\{ u \sum_{m=1}^M y_m N_m(t) \right\} \cdot \prod_{m=1}^M e^{(\lambda_m - \tilde{\lambda}_m)t} \left( \frac{\tilde{\lambda}_m}{\lambda_m} \right)^{N_m(t)} \right] \\
 &= \prod_{m=1}^M \exp\{(\lambda_m - \tilde{\lambda}_m)t\} \cdot \mathbb{E} \left[ \exp \left\{ \left( u y_m + \log \frac{\tilde{\lambda}_m}{\lambda_m} \right) N_m(t) \right\} \right] \\
 &\stackrel{\textcircled{3}}{=} \prod_{m=1}^M \exp\{(\lambda_m - \tilde{\lambda}_m)t\} \exp \left\{ \lambda_m t \left( e^{u y_m + \log(\tilde{\lambda}_m/\lambda_m)} - 1 \right) \right\} \\
 &= \prod_{m=1}^M \exp \left\{ (\lambda_m - \tilde{\lambda}_m)t + \tilde{\lambda}_m t e^{u y_m} - \lambda_m t \right\} \\
 &= \prod_{m=1}^M \exp \left\{ \tilde{\lambda}_m t (e^{u y_m} - 1) \right\} \\
 &\stackrel{\textcircled{4}}{=} \prod_{m=1}^M \exp \left\{ \tilde{\lambda}_m t \tilde{p}(y_m) e^{u y_m} - \tilde{\lambda}_m t \right\} \\
 &\stackrel{\textcircled{5}}{=} \exp \left\{ \tilde{\lambda} t \left( \sum_{m=1}^M \tilde{p}(y_m) e^{u y_m} - 1 \right) \right\}.
 \end{aligned}$$

According to (11.3.5), this is the moment-generating function for a compound Poisson process with intensity  $\tilde{\lambda}$  and jump-size distribution (11.6.14).  $\square$

**Lemma 5.2.1.** Let  $t$  satisfying  $0 \leq t \leq T$  be given and let  $Y$  be an  $\mathcal{F}(t)$ -measurable random variable. Then

$$\tilde{\mathbb{E}} Y = \mathbb{E}[Y Z(t)]. \quad (5.2.8)$$

$$\varphi_{N(t)}(u) = \mathbb{E} e^{uN(t)} = \exp\{\lambda t(e^u - 1)\}. \quad (11.3.4)$$

$\textcircled{1}$  By Lemma 5.2.1

$$\textcircled{2} \quad Z_m(t) = e^{(\lambda_m - \tilde{\lambda}_m)t} \left( \frac{\tilde{\lambda}_m}{\lambda_m} \right)^{N_m(t)} \quad \text{and} \quad Z(t) = \prod_{m=1}^M Z_m(t).$$

$$Q(t) = \sum_{m=1}^M y_m N_m(t)$$

$\textcircled{3}$  By 11.3.4

$$\textcircled{4} \quad \tilde{\mathbb{P}}\{Y_i = y_m\} = \tilde{p}(y_m) = \frac{\tilde{\lambda}_m}{\tilde{\lambda}}. \quad (11.6.14)$$

$$\Rightarrow \tilde{\lambda}_m = \tilde{P}(y_m) \cdot \tilde{\lambda}$$

$$\textcircled{5} \quad \tilde{\lambda} = \sum_{m=1}^M \tilde{\lambda}_m$$

The Radon-Nikodým derivative process  $Z(t)$  of (11.6.11) may be written as

$$Z(t) = \exp \left\{ \sum_{m=1}^M (\lambda_m - \tilde{\lambda}_m) t \right\} \cdot \prod_{m=1}^M \left( \frac{\tilde{\lambda} \tilde{p}(y_m)}{\lambda p(y_m)} \right)^{N_m(t)} = e^{(\lambda - \tilde{\lambda})t} \prod_{i=1}^{N(t)} \frac{\tilde{\lambda} \tilde{p}(Y_i)}{\lambda p(Y_i)}.$$

This suggests that if  $Y_1, Y_2, \dots$  are not discrete but instead have a common density  $f(y)$ , then we could change the measure so that  $Q(t)$  has intensity  $\tilde{\lambda}$  and  $Y_1, Y_2, \dots$  have a different density  $\tilde{f}(y)$  by using the Radon-Nikodým derivative process

$$Z(t) = e^{(\lambda - \tilde{\lambda})t} \prod_{i=1}^{N(t)} \frac{\tilde{\lambda} \tilde{f}(Y_i)}{\lambda f(Y_i)}. \quad (11.6.15)$$

This is in fact the case, although the proof, given below, is harder than the one just given for the case of a discrete jump-size distribution.

To avoid division by zero in (11.6.15), we assume that  $\tilde{f}(y) = 0$  whenever  $f(y) = 0$ . This means that if a certain set of jump sizes has probability zero under  $\mathbb{P}$ , then it will also have probability zero under  $\tilde{\mathbb{P}}$  considered in Theorem 11.6.7 below.

$$\begin{aligned} Z(t) &= \prod_{m=1}^M e^{(\lambda_m - \tilde{\lambda}_m)t} \left( \frac{\tilde{\lambda}_m}{\lambda_m} \right)^{N_m(t)} \\ &= \exp \left\{ \sum_{m=1}^M (\lambda_m - \tilde{\lambda}_m) t \right\} \cdot \prod_{m=1}^M \left( \frac{\tilde{\lambda} \tilde{p}(y_m)}{\lambda p(y_m)} \right)^{N_m(t)} \\ &\stackrel{\text{換次數}}{\rightarrow} = \exp \{ (\lambda - \tilde{\lambda}) t \} \cdot \prod_{i=1}^{N(t)} \frac{\tilde{\lambda} \tilde{p}(Y_i)}{\lambda p(Y_i)} \\ &= \exp \{ (\lambda - \tilde{\lambda}) t \} \cdot \prod_{i=1}^{N(t)} \frac{\tilde{\lambda} \tilde{f}(Y_i)}{\lambda f(Y_i)} \end{aligned}$$

discrete 換  
continuous